

Quantitative Techniques – Term 2

Laboratory 2

26 January 2006

Overview

The objectives of this lab are:

- Apply known statistical concepts (correlation, tests with two samples) to competition policy;
- Guide you through an overview of other statistical tests of prices and price trends.

Start Stata 9 from the Windows Start menu (Programs, F. Departmental Software, E. Social Science, Stata). Do not update the version that is installed on your machine.

Before starting your tasks, make sure that the following windows are visible on your screen and choose their size so that they do not overlap: Review, Variables, Results, Command.

Task 1 – From last week: import a data file¹

- Go to www.staff.city.ac.uk/a.banal-estanol/teaching.htm and download Lab2Prices.csv to your directory.
Note that the names of the variables are in the first row, immediately followed by your data.

The file contains time series of weekly prices of retail gasoline prices in \$cents per gallon in different areas of the United States:²

- East Coast;³
- Midwest;⁴
- Gulf Coast;⁵
- Rocky Mountains;⁶
- West Coast.⁷

In addition, the dates when prices are measured are also included, as is a variable “Time”, which simply counts the number of weeks.

- Type in the Command window:
insheet using “*your directory*\Lab2Prices.csv”

¹ Reference: <http://www.stata.com/help.cgi?infilng>

² The source of the data is <http://www.economagic.com/doewkly.htm>

³ CT, MA, VT, NH, ME, RI, NY, NJ, PA, DC, MD, DE, WV, VA, NC, SC, GA, FL.

⁴ MI, OH, KY, TN, IN, WI, IL, MN, IO, MO, ND, SD, NE, KS, OK.

⁵ AL, MS, AR, LA, TX, NM.

⁶ MT, WY, CO, ID, UT.

⁷ AK, WA, OR, NV, CA, AZ, HI.

This command loads the data in Stata. Note that Stata recognizes commands in lower case letters. Use “” when you specify a filename.

Alternatively, you can use the Menu bar: File, Import, ASCII data created by a spreadsheet.

- Save the file as Lab2Prices.dta. This is the extension used by Stata for datasets.

Task 2 – Cross-sectional price tests: independent samples⁸

This task builds on the statistical concepts you have learned in the first term. We want to test the null hypothesis that the average price in Midwest (μ_M) is equal to the average price on the East Coast (μ_E) (e.g. $H_0: \mu_M = \mu_E$). Stata allows you to test this hypothesis against three alternative hypotheses at the same time.

You can:

- Go to the Command window and type **ttest eastcoast== midwest, unpaired**; OR
- From the Menu bar, go to Statistics... Summaries, tables and tests... Classical tests of hypotheses... Two-sample mean comparison test. Then choose the variables and tick “unpaired”.

```
. ttest eastcoast== midwest, unpaired
```

Two-sample t test with equal variances

Variable	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
eastcoast	96	111.1375	.1012666	.9922064	110.9365	111.3385
midwest	96	109.001	.1331427	1.304526	108.7367	109.2654
combined	192	110.0693	.1137245	1.575813	109.845	110.2936
diff		2.136458	.1672779		1.806498	2.466419

Degrees of freedom: 190

Ho: mean(eastcoast) - mean(midwest) = diff = 0

Ha: diff < 0	Ha: diff != 0	Ha: diff > 0
t = 12.7719	t = 12.7719	t = 12.7719
P > t = 1.0000	P > t = 0.0000	P > t = 0.0000

Task 3 – Hedonic price analysis⁹

This exercise is based on Exercise 3, Chapter 4 of Berndt, E.R. (1991), “The Practice of Econometrics: Classic and Contemporary”, Addison-Wesley.

The purpose of the exercise is to construct a price index for computers using hedonic regression techniques.

- Download Lab2Hedonic.dta from www.staff.city.ac.uk/a.banal-estanol/teaching.htm

⁸ <http://www.stata.com/help.cgi?ttest>

⁹ Reference on the use of hedonic price indexes in practice: <http://www.bls.gov/ppi/ppicomqa.htm>

Open Stata in a new window and then you can open the file by either:

- Open – using the Menu bar or the icon on the toolbar; or
- Type in the command window:

use “*your directory*\Lab2Hedonic.dta”

Again, note the use of “ “ when you specify a filename.

The file contains the following variables:

- Year: year in which the model was introduced (from 1954 to 1965);
- Lnrent: natural log of monthly rental of computers;
- Lnmult: natural log of average time to obtain and complete multiplication instructions;
- Lnaccess: natural log of average time to access information from memory;
- Lnmem: natural log of memory size;
- D61: a variable which takes value 1 when year=61 and zero otherwise;
- D62, d63, d64, d65: constructed similarly as for d61.

By using OLS, estimate the parameters of the equation for years '60 to '65:¹⁰

$$\text{Lnrent} = \text{constant} + a*d61 + b*d62 + c*d63 + d*d64 + e*d65 + f*\text{lnmult} + g*\text{lnaccess} + h*\text{lnmem} + \text{error}$$

(You can also click on the variables in the left-hand side Variables window.)

```
. reg lnrent d61 d62 d63 d64 d65 lnmult lnaccess lnmem if year>59
```

Source	SS	df	MS	Number of obs = 82		
Model	106.679895	8	13.3349869	F(8, 73)	=	88.77
Residual	10.965562	73	.150213178	Prob > F	=	0.0000
-----				R-squared	=	0.9068
-----				Adj R-squared	=	0.8966
Total	117.645457	81	1.45241305	Root MSE	=	.38757

lnrent	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
d61	-.1386585	.1679443	-0.83	0.412	-.473371	.1960541
d62	-.4916432	.175347	-2.80	0.006	-.8411092	-.1421771
d63	-.5948134	.1676572	-3.55	0.001	-.9289537	-.2606731
d64	-.9265809	.1678509	-5.52	0.000	-1.261107	-.5920546
d65	-1.14627	.1671586	-6.86	0.000	-1.479416	-.8131232
lnmult	-.0711882	.0282091	-2.52	0.014	-.1274089	-.0149675
lnaccess	-.1349376	.0292385	-4.62	0.000	-.1932099	-.0766654
lnmem	.5770442	.0357484	16.14	0.000	.5057978	.6482907
_cons	-.0733124	.3176761	-0.23	0.818	-.7064399	.5598152

- Note that the coefficients highlighted above can be interpreted as elasticities.
- The coefficients on d61, d62, d63, d64 and d65 can be used to construct quality-adjusted price indices with respect to 1960, as follows:

¹⁰ <http://www.stata.com/help.cgi?regress>

Year	Coefficient (e.g. α)	Price index - antilogarithm (e.g. e^α)	% Change with respect to 1960
1961	-.1386	0.871	-12.9%
1962	-.49164	0.612	-38.8%
1963	-.5948	0.552	-44.8%
1964	-.9265	0.396	-60.4%
1965	-1.1462	0.318	-68.2%

Task 4 – Price correlation¹¹

- Using the same data as in Task 2, calculate the correlation between the prices in the different geographic areas. (Remember you can also use the Menu bar: Statistics... Summaries, tables and tests... Summary statistics... Correlations and covariances).

```
. correlate eastcoast midwest westcoast gulfcoast rocky
(obs=96)

-----+-----
| eastco~t  midwest westco~t  gulfco~t   rocky
-----+-----
eastcoast |    1.0000
midwest   |    0.6658    1.0000
westcoast |    0.2490    0.5745    1.0000
gulfcoast |    0.7134    0.9225    0.6322    1.0000
rocky     |    0.1291    0.6182    0.8434    0.6280    1.0000
```

- Repeat the correlation analysis for the log of prices (LECG report, page 54).
Reminder: to generate a new variable, you need the command “generate”. E.g. **gen lneastcost=ln(eastcoast)**

```
. correlate lneastcost lnmidwest lnwestcoast lngulfcoast lnrocky
(obs=96)

-----+-----
| lneast~t lnmidw~t lnwest~t lngulf~t  lnrocky
-----+-----
lneastcost |    1.0000
lnmidwest  |    0.6647    1.0000
lnwestcoast |    0.2469    0.5761    1.0000
lngulfcoast |    0.7122    0.9224    0.6335    1.0000
lnrocky    |    0.1340    0.6231    0.8451    0.6340    1.0000
```

- Repeat the analysis for the differences in the logarithms of prices (approximation of growth rates).
- In order to do this, we will use the difference operator. First, you need to declare your dataset to be time-series by typing **tsset time**

```
tsset time
time variable:  time, 1 to 96
```

- The difference operator is written **D.** or **d.** followed by the name of the variable. You don't need to generate a variable separately, you can simply use the command you need followed by the difference operator.

¹¹ <http://www.stata.com/help.cgi?corr>

```
. correlate D.lneastcoast D.lnmidwest D.lnwestcoast D.lngulfcoast D.lnrocky
(obs=95)
```

	D.	D.	D.	D.	D.
	lneast~t	lnmidw~t	lnwest~t	lngulf~t	lnrocky
lneastcoast	1.0000				
lnmidwest	0.6454	1.0000			
lnwestcoast	0.2513	0.2297	1.0000		
lngulfcoast	0.5669	0.6636	0.4231	1.0000	
lnrocky	0.1064	0.2458	0.3826	0.2795	1.0000

Task 5 – Granger causality¹²

- Using the same data as in Task 2, we carry out a bivariate Granger causality test, assuming a lag length 3;
- Note that, for simplicity, we are not removing the effects of any common factors except for the time trend, to demonstrate the syntax of the command;
- The first step is to estimate the following model by OLS (from LECG, 1999, page 59).¹³

$$P_1 = \sum_{s=1}^T \beta_s P_{1t-s} + \sum_{s=1}^T \gamma_s P_{2t-s} + u_t$$

- In our example, P1 is the price on the East Coast and P2 the price on the West Coast. As we have set the lag length = 3, we have that T = 3.
- In order to simplify the notation we use the lag operator, **L**, or **I**, followed by the name of the variable.

```
. reg eastcoast L(1/3).eastcoast L(1/3).westcoast time
```

Source	SS	df	MS	Number of obs =	93
Model	79.9069453	7	11.4152779	F(7, 85) =	82.76
Residual	11.7248659	85	.137939599	Prob > F =	0.0000
				R-squared =	0.8720
				Adj R-squared =	0.8615
Total	91.6318112	92	.995997947	Root MSE =	.3714

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
eastcoast					
L1.	1.044364	.1066324	9.79	0.000	.8323503 1.256378
L2.	.1012993	.1501328	0.67	0.502	-.1972051 .3998036
L3.	-.3537807	.1092032	-3.24	0.002	-.570906 -.1366555
westcoast					
L1.	.0362912	.0667641	0.54	0.588	-.0964537 .1690361
L2.	-.0004602	.1021941	-0.00	0.996	-.2036495 .2027291

¹² There is an alternative way of testing for Granger causality directly, but it requires installing an additional routine from the web. For this reason, it is not feasible in the University computer room. The description of how to install this package follows, for your information. Type **net install gcause.pkg** in the Command window. Alternatively, from the Menu bar, Help... Search... Search all: the keyword to search for is “gcause”. Then follow the instructions to install the package.

¹³ References for Granger causality test: Greene, W. (2003), “Econometric Analysis”, Prentice Hall; Hamilton, J.D. (1994), “Time Series Analysis”, Princeton University Press.

L3.		.0131382	.0676851	0.19	0.847	-.121438	.1477144
time		-.0038358	.0023731	-1.62	0.110	-.0085542	.0008826
_cons		17.31988	5.844017	2.96	0.004	5.700412	28.93936

- Then we perform an F-test for causality, i.e. the null hypothesis is that the coefficients on P2 are jointly zero.

```
. test 1.westcoast 12.westcoast 13.westcoast

( 1) L.westcoast = 0
( 2) L2.westcoast = 0
( 3) L3.westcoast = 0

F( 3, 85) = 1.55
Prob > F = 0.2070
```

Task 6 – Dynamic price regressions

Using the same data as in Task 2, we need to test whether the time series of prices are stationary or to identify their order of integration, then verify if they are co-integrated. This task is at a very introductory level and has the only objective of illustrating the steps required in the estimation of dynamic regressions.

1. Dickey-Fuller tests for unit roots.¹⁴

Consider once again prices on the East Coast (P1) and on the West Coast (P2). In this example, we are going to use `lneastcoast` and `lnwestcoast`.

The null hypothesis is that the series is non-stationary.

You can run the test in Stata by clicking on the Menu bar... Statistics... Time Series... Tests... Augmented Dickey Fuller unit-root test.

- Choose the variable you want to run the test for (e.g. `lneastcoast`);
- Tick “Display regression table”;
- Set the maximum number of lags to zero (setting the lags > 0 means that you are running an “Augmented Dickey Fuller” test – we do not consider this for the moment).

¹⁴ This is only one of the available tests for unit roots in time-series data. In our example, we are testing whether the series is $x_t = a + x_{t-1} + e_t$
References: Greene, op. cit., Section 20.3.4; Charemza, W., and D. Deadman (1992), “New Directions in Econometric Practice”, Edward Elgar.
<http://www.stata.com/help.cgi?dfuller>

```
. dfuller lneastcoast, regress lags(0)

Dickey-Fuller test for unit root                                Number of obs   =           95

----- Interpolated Dickey-Fuller -----
                Test          1% Critical   5% Critical   10% Critical
                Statistic      Value         Value         Value
-----
Z(t)           -1.666          -3.517        -2.894        -2.582
-----
MacKinnon approximate p-value for Z(t) = 0.4487
```

```
-----
D.
lneastcoast |          Coef.   Std. Err.    t    P>|t|    [95% Conf. Interval]
-----+-----
lneastcoast |
   L1       |  -.0724415   .0434823   -1.67  0.099   - .1587886   .0139056
_cons      |   .3413512   .2048254    1.67  0.099   - .0653915   .7480938
-----
```

The test statistic (-1.666) is larger (less negative) than the critical values and therefore we cannot reject the null hypothesis of non-stationarity. We use again the difference operator and we repeat the test on the differenced series, i.e. $lneastcoast_t - lneastcoast_{t-1}$.

```
. dfuller d.lneastcoast, regress lags(0)

Dickey-Fuller test for unit root                                Number of obs   =           94

----- Interpolated Dickey-Fuller -----
                Test          1% Critical   5% Critical   10% Critical
                Statistic      Value         Value         Value
-----
Z(t)           -7.519          -3.518        -2.895        -2.582
-----
MacKinnon approximate p-value for Z(t) = 0.0000
```

```
-----
D2.
lneastcoast |          Coef.   Std. Err.    t    P>|t|    [95% Conf. Interval]
-----+-----
lneastcoast |
   LD       |  -.7929278   .1054501   -7.52  0.000   -1.002361   -.5834947
_cons      |   .0001496   .0003803    0.39  0.695   - .0006056   .0009049
-----
```

We can now reject the null hypothesis that $(lneastcoast_t - lneastcoast_{t-1})$ is non-stationary. This leads us to the conclusion that $lneastcoast$ is a series integrated of order one.

Repeating the steps above for $lnwestcoast$, we obtain analogous results for $lnwestcoast$. Both series are then integrated of the same order and we can proceed to the next step.

2. Co-integration test.

This test involves verifying whether there exists a linear combination of the two series that is stationary (see LECG, page 64). We therefore need to test whether the fitted residuals from the linear model $lneastcoast_t = a + b*lnwestcoast_t + e_t$ are stationary.

- Estimate the following model

```
. reg lneastcoast lnwestcoast
```

Source	SS	df	MS	Number of obs = 96		
Model	.000462415	1	.000462415	F(1, 94)	=	6.10
Residual	.007120168	94	.000075746	Prob > F	=	0.0153
				R-squared	=	0.0610
				Adj R-squared	=	0.0510
				Root MSE	=	.0087
Total	.007582583	95	.000079817			

lneastcoast	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
lnwestcoast	.1073602	.0434519	2.47	0.015	.0210855	.193635
_cons	4.194378	.2089846	20.07	0.000	3.779434	4.609322

- Calculate the residuals

```
. predict resid, residuals
```

- Test for stationarity of the residuals

In order to reject the null hypothesis of non-stationarity, the test statistic has to be more negative than the critical value for this test. At a 5% level, for around 100 observations, the critical value is -4.14 .

Note that there are separate tables for stationarity tests when applied to the residuals. For this reason, the critical value does not correspond to that given by Stata. (For an explanation see Charemza and Deadman, page 129).

```
. dfuller resid, regress lags(0)
```

```
Dickey-Fuller test for unit root
```

Test Statistic	Interpolated Dickey-Fuller		
	1% Critical Value	5% Critical Value	10% Critical Value
Z(t)	-3.517	-2.894	-2.582

MacKinnon approximate p-value for Z(t) = 0.3559

D.resid	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
resid						
L1	-.0800248	.0432527	-1.85	0.067	-.1659161	.0058664
_cons	.0000516	.0003689	0.14	0.889	-.0006809	.0007841

As the test statistic is less negative than the critical value, we conclude that the residuals are not stationary and therefore the two variables are not cointegrated.

Answer to question in Lab 1, related to Task 5:

How to test the following hypothesis?

H_0 : coefficient on educ = 0.1, coefficient on kidsge6 = -0.03

```
. test (educ=0.1) (kidsge6=-0.03)
```

```
( 1) educ = .1
( 2) kidsge6 = -.03

F( 2, 421) = 0.27
Prob > F = 0.7639
```